FARMLAND WITH SAMPLE AREA SEGMENTS IN AGRICULTURAL SURVEYS.*

W. A. Hendricks, D. T. Searls and D. G. Horvitz Research Triangle Institute

1. Introduction

The application of area sampling techniques to sample farm surveys requires some rule for associating farms and farmland with the selected area segments. The rule that was adopted in the United States about a quarter of a century ago, when area sampling first began to be applied to farm surveys in this country, and which has been used most generally up to the present time, is the so-called "headquarters rule." By this rule a point on every farm which can be defined rigorously, and which can be identified by interviewers in the field, is employed as a reference point. If this point for a particular farm falls within the boundaries of the area segment, the farm is regarded as being "in" the segment. But if the reference point is outside the segment, the farm is considered outside the segment, even though some or most of all land in the farm may fall inside the segment.

The Census Bureau and other agencies involved in farm surveys have developed elaborate definitions of the reference point which is to serve as the "headquarters" of the farm in any particular instance, such as the following:

- a. If the operator of the farm lives on the farm, his house is the headquarters.
- b. If the operator does not live on the farm but there is one and only one house on the farm, that house is the headquarters.
- c. If there is more than one house on the farm and the operator does not live on the farm, the house of greatest value is the headquarters.
- d. If there are no houses on the farm but other buildings are present, the building of greatest value is the headquarters.
- e. If there are no buildings on the farm, the "main entrance" to the farm is the headquarters.
- f. If no point can be identified as the main entrance, the corner of the farm farthest west and farthest north (in that order) is the headquarters.

Such sequences are established to permit the use of reference points that can be identified by interviewers with the least difficulty and possibility of error. However, errors in associating farms with area segments by a "headquarters" rule are usually all to prevalent. The problem is aggravated by difficulties and errors that often arise in deciding which parcels of land constitute the "farm."

To avoid these troublesome problems, the <u>agency of the</u> United States Department of Agricul-

"This study was conducted for the Bureau of the Census, United States Department of Commerce, under Contract No. CcO-7574. ture now known as the Statistical Reporting Service has in recent years made extensive use of a so-called "Closed Segment" rule. By this rule interviewers must account only for items on land that lies entirely within the segment boundaries. To obtain estimates of numbers of farms, all persons, or a sample of them, living within the segment boundaries must be screened to determine how many are farm operators. This is accomplished by ascertaining the nature and extent of their agricultural operations, if any, regardless of where they are performed. These persons are also interviewed to obtain data that can only be obtained for the farm as a whole, such as sales and purchases of various commodities. Wherever possible, estimates of such items are made to conform to related information obtained from land within the segment boundaries. For example, universe estimates of cattle inventories are obtained from the numbers found within the sample segments on the date of the enumeration. But estimates of sales are obtained by applying the ratio of farm sales to farm inventories, for entire farms as reported by farm operators in the segments, to the estimate of total inventories derived only from numbers within the segment boundaries.

This "closed segment" rule has several advantages over a "headquarters" rule, but it also has some disadvantages. On the plus side we note (1) the rule is less troublesome for interviewers to apply in the field, (2) as the land to be accounted for in each sample segment is depicted on the interviewers' aerial photos, both interviewers and supervisors can recognize gross errors in reported data that might otherwise pass undetected, (3) reporting errors can be reduced because respondents are asked to report separately for specific tracts of land pointed out to them on the aerial photos, and in many cases where their holdings inside the segment boundaries represent only a portion of all their holdings, they are not required to disclose information about the portions outside the segment, and finally (4) between-segment sampling variation can be reduced because the boundaries of each segment place a limit on the total land to be accounted for in connection with that segment.

On the minus side we observe that farm operators who live in cities and towns can only be counted if some sample segments are allocated to urban areas. The task of identifying farm operators in such areas is often formidable and considerable undercounting may result. Furthermore, there is also the risk that some urban residents classified as farm operators may have tenants on their farms who would be classified as the operators if they fell into the opencountry portion of the sample. This is perhaps the greatest defect in the closed segment procedure. As pointed out previously, the fact that the method is not well adapted to surveys seeking data which can be reported conveniently only for farms as a whole must also be considered on the debit side.

Another approach, which for want of a better name has been called the "weighted segment" procedure, seems to offer a solution that retains many of the advantages of both the "headquarters" and "closed segment" rules and is also free of the most serious objections levelled at those two methods. So far as the authors of this paper are aware, it has not been applied to farm surveys in the United States as yet, although an agency of the Department of Agriculture has sponsored a rural land-ownership survey in which this method was employed. The Statistical Reporting Service made use of a similar principle in a survey for estimating the availability of farm grain storage facilities a few years ago.

As applied to farm surveys, the "weighted segment" approach regards every farm with some of its land inside a sample segment as associated with that segment. Data are recorded for every such farm as a whole, but are multiplied by the fraction of the farmland in the farm that falls within the segment before incorporating them into segment totals.

This procedure has a number of advantages over the two previous rules. First, it is a rule that can be applied by interviewers in the field with less difficulty and possibility of error than a headquarters rule. Also the need for canvassing urban areas is eliminated and the sample can be confined to the open country. All data are recorded for entire farms; hence no special treatment is required for items that can not be reported conveniently for portions of farms as in the closed segment approach. At the same time the weighting of the data for each farm by the fraction of its land falling within. the segment can reduce the between-segment variability of segment totals to a level comparable to that obtained with closed segment data. In fact, this variability can be expected to be lower because of the averaging effect achieved by prorating entire-farm data to land within the segment, rather than recording data only for the land within the segment.

Among the disadvantages, as compared with the closed segment procedure, we observe that interviewers are still faced with the problem of deciding which parcels of land must be defined as a "farm" and this is sometimes difficult. However, errors introduced by including too much land, or too little, in defining a farm tend to be partially neutralized by the weighting procedure. The fact that all land reported as being in a farm may not be covered by the interviewers' aerial photo eliminates some of the visual verification that can be performed by supervisors on closed segment data. The weighting that must be applied to individual farm data is a computational step that is not required with either the headquarters or closed segment rules, but with automatic data processing procedures that are now in rather general use this is not a serious matter. With sample segments of a given size, data must be recorded for about twice as many farms by this rule as compared with a headquarters rule. With long questionnaires this can increase the time required to be spent in each sample segment, although it would by no means double the time required with a headquarters rule. The proper application of a headquarters rule requires a complete canvass of each sample segment to ascertain the places eligible for enumeration; considerable time is often consumed in screening out ineligible places.

2. Objectives of the Present Study

The purpose of the present study was to investigate the sampling variances encountered with each of the preceding rules for associating farms with sample area segments. Although some fragmentary information on the headquarters and closed segment rules has accumulated over the past few years, no systematic comparison of the two approaches in this respect has ever been made. So far as the weighted segment rule is concerned, objective data on variances are nonexistent.

The data employed in this study were obtained in the 1954 and 1959 Census of Agriculture Evaluation Programs. Data from 384 identical segments in 175 primary sampling units (PSU) were available for both of those years in a form that made such an analysis possible. In addition to detailed information about the characteristics of each farm covered in the Evaluation Program, data were recorded separately for the portions of those farms that fell inside the segment boundaries. To investigate the behavior of the headquarters rule, only two alternative reference points on each farm were considered as headquarters: (1) if the operator lived on the farm, his residence was the headquarters, and (2) if the operator did not live on the farm, the point on the farm farthest west and farthest north (in that order) was the headquarters. Sketches of the farm in relation to the segment boundaries were available to make the appropriate determination in each case.

All relevant information was placed on punch cards to facilitate the computations, which were performed on an IEM 1401 computer. The items studied in the analysis were:

- Numbers of farms (omitted in closed segment analysis)
- 2. Acres of farmland
- 3. Acres of cropland harvested
- 4. Acres of corn harvested
- 5. Acres of wheat harvested
- 6. Acres of cotton harvested
- 7. Acres of soybeans harvested
- 8. Acres of oats harvested
- 9. Acres of hay cut
- 10. Acres of tobacco harvested

Between segment variability within primary units was estimated for the above items separately for each of three regional strata of the United States and for each of the three rules employed to associate sample farms with sample area segments. Coefficients of correlation between 1954 and 1959 data were also computed for each rule of association. The variances and covariances within primary units were computed under the assumption that a large sample of farms in the United States would be a one-stage design with sample segments allocated proportionally to all PSU's in the universe rather than a two-stage design of the kind used in the Evaluation Program. To indicate how much improvement could be attained by basing 1959 estimates of agricultural items upon 1954 data through the use of difference, ratio, and regression estimators, the variability of such estimates was computed on a per-segment basis for comparison with the per-segment variances of the 1959 segment totals. To investigate possible gains in precision by excluding large farms from the area sample, all variances and covariances were computed with data for farms of 2,000 acres or more included and excluded.

The three regional strata and the number of sample segments in each are as follows:

Region I- North (154 segs.)	Region II- South (116 segs.)	Region III- West (114 segs.)
(154 segs.) Connecticut Illinois Indiana Iowa Maine Massachusetts Michigan Minnesota Missouri New Hampshire New Jersey New York Ohio Pennsylvania Rhode Island	(116 segs.) Alabama Arkansas Delaware Florida Georgia Kentucky Louisiana Maryland Mississippi N. Carolina S. Carolina Tennessee Virginia	(114 segs.) Arizona California Colorado Idaho Kansas Montana Nebraska Nevada New Mexico N. Dakota Oklahoma Oregon S. Dakota Texas Utah
Vermont West Virginia Wisconsin		Washington Wyoming

These regions do not conform exactly to the regions for which official agricultural statistics are customarily summarized. The regional strata employed here were adopted mainly to achieve greater equality in the number of sample segments per stratum.

3. Computational Methods

The 175 primary units in the three regions had been selected in 1954 with unequal probabilities and the sampling rates within those PSU's had been adjusted to arrive at a self weighted sample. For the present analysis some PSU's containing only one sample segment each needed to be combined with other PSU's to permit the computation of within-PSU variability. After those combinations were made, the 384 segments were contained in 124 new PSU's which were regarded as being selected with probabilities comparable to the original PSU. The within-PSU sampling rates were regarded as proportional to the reciprocals of those probabilities. The average 1959 variance between segments within PSU's for a given region was therefore computed from the formula

$$s_{y}^{2} = \frac{1}{n} \sum_{i} \frac{n_{i}}{n_{i}-1} \sum_{j} (y_{ij} - \bar{y}_{i})^{2}$$
(1)

in which

- y_{ij} = a 1959 segment total for the j-th segment in the i-th PSU.
- y_i = the per-segment average of the y_{ij} for all segments in the i-th PSU.
- n_i = the number of sample segments in the i-th PSU.
- n = the total number of sample segments
 in the region.

Average covariances between 1954 and 1959 data were computed in a similar fashion:

$$s_{xy} = \frac{1}{n} \sum_{i} \frac{n_{i}}{n_{i}-1} \sum_{j} (x_{ij} - \bar{x}_{i})(y_{ij} - \bar{y}_{i}) \quad (2)$$

in which \mathbf{x}_{ij} and \mathbf{y}_{ij} are comparable segment totals in 1954 and ij 1959.

Clearly, if an estimated universe total of an agricultural item is to be made for a current year, such as 1959, without reference to related data for previous years, the relvariance (RV) of that estimate will be equal to the relvariance of the per-segment average for that item in the sample. If such an estimate is represented by Y_1 , we have

$$RV(Y_1) = \frac{s_y^2}{n\bar{y}^2}$$
(3)

in which s^2 is the between segment variance of the sample^y segment totals, \bar{y} is the per-segment average of all segment totals, and n is the number of segments in the sample.

If a difference estimate, which makes use of data for the universe and the sample in a previous year such as 1954 is computed, such an estimate takes the form

$$Y_2 = X + (Y_1 - X_1)$$
 (4)

in which X is the universe total in the base year, Y_1 is the current year estimate obtained by applying the reciprocal of the sampling rate to the current year sample total, and X_1 is the corresponding estimate of X derived from base year data in the sample. The relvariance of Y_2 is given by

$$RV(Y_2) = \frac{s_x^2 + s_y^2 - 2s_{xy}}{n\bar{y}^2}$$
(5)

in which the various terms are self-explanatory.

Instead of employing a difference estimator, one might invoke a ratio estimate of the form

$$Y_3 = \frac{Y_1}{X_1} X \tag{6}$$

The relvariance of Y_3 , written in terms of the relvariances of X_1 and Y_1 , together with the relative covariance (RCV) of X_1 and Y_1 is approximately:

$$RV(Y_3) = RV(X_1) + RV(Y_1) - 2RCV(X_1Y_1)$$
 (7)

A third alternative would be to employ a regression estimator of the form

$$Y_4 = Y_1 + b(X - X_1)$$
 (8)

in which b is the estimated regression coefficient of y on x and the other symbols have the same meaning as before. The relvariance of Y_{I_1} is approximately

$$RV(Y_4) = (1 - r_{xy}^2) RV(Y_1)$$
 (9)

in which r is the average within PSU coefficient of xy correlation between 1954 and 1959 totals for the same segments.

4. Numerical Results •

Simple Expansion Estimates

The relvariance of a simple expansion estimate, as indicated previously, is identical with the relvariance of the per-segment average of that item. The estimated relvariances for all farms for each of the selected items under consideration are shown in Table 1 on a per-segment basis (n = 1) by region for each of the three association rules, and also with farms of 2,000 acres or more excluded. No such large farms were in the Region I sample.

The table indicates that the exclusion of large farms from the sample (Regions II and III) did not reduce the relative variances to any great degree. Aside from farmland itself, some reduction was effected in the relative variances of a few items such as corn, oats, and hay acreages in Region II.

The Closed Segment and Weighted Segment Rules tend to yield the lowest variances, with the Weighted Segment showing a slight edge over the Closed Segment. Table 2 shows that the use of the Weighted Segment Rule about doubles the number of farms from which information is obtained as compared with the Headquarters approach. This does not double the interview cost because interviewers must account for all land in a segment to identify farms with headquarters in the segment.

Table 2. Farms with Land in Segment by Location of Headquarters, 1959 EPA Rural Area Sample

	·Number	of farms in sam	ple
Region ¹ /	With Hq. in Segment	With Hq. Not in Segment	Total
North South West	1,192 1,072 291	1,251 1,033 240	2,443 2,105 531
U. S.	2,555	2,524	5,079

 $\frac{1}{7}$ The regions are as defined for the 1959 Census, not as for this study.

1954-59 Correlation Coefficients

Average correlation coefficients between segments within PSU are shown in Table 3 for the items covered in Table 1. Correlations are highest in Region I and lowest in Region III. In Region II they are a bit larger when large farms are retained in the sample, but in Region III the reverse seems to be true in several instances. One extremely large place of about 128,000 acres in Region III was omitted from the computations because it tended to dominate the results unduly. Generally speaking, correlations are highest with the Closed and Weighted Segment Rules of association. Of these two, the difference is again in favor of the Weighted Segment Rule.

The correlations are of sufficient magnitude, almost in general, to yield a considerable increase in statistical efficiency with estimation procedures which include prior Census or base year information as compared to simple expansion estimation. This is verified in Tables 4, 5 and 6. The average correlations between segments within PSU's for the years immediately following a Census year should be even larger than those obtained for the 5-year interval in this study. It should be pointed out, however, that base year data assembled for the sample segments during a Census should be in a form comparable to the data that will be collected according to the specified association rule in intercensal sample surveys.

The negative correlations for cotton acreage in Region II with the Headquarters Rule were due to two segments with fairly large acreages in 1954 but small cotton acreages, according to this rule in 1959. This was due to a shift of the headquarters of one or two farms out of these segments in 1959, rather than a larger reduction in cotton acreage between 1954 and 1959. The correlations for cotton acreage in this region for the Closed and Weighted Segment Rules remained substantially postive.

Difference Estimates

The relative variances of 1959 difference estimates, based on 1954 data from the same segTable 1. Estimated Average Within Primary Sampling Unit Relative Variances (n=1) for <u>Simple Expansion Estimates</u> obtained with Three Rules of Associating Farm Land with Sample Area Segments, by Region, with and without Farms of 2000 or More Acres.

		All Farms	1	Excluding Large Farms			
	Asso	Association Rule			Association Rule		
Item	Head-	Closed	Weighted	Head-	Closed	Weighted	
	quarters	Segment	Segment	quarters	Segment	Segment	
			Reg	ion I			
Farms	0.388	*	0.381	0,388	*	0.381	
Farmland	• 544	•420	•420	• 544	. 420	. 420	
Cropland Harvested	.720	.561	. 567	•720	.561	•567	
Corn Acreage	.913	. 858	.746	.913	. 858	•746	
Wheat Acreage	2.150	2.625	1.998	2.150	2.625	1.998	
Cotton Acreage	**	**	**	**	**	**	
Soybean Acreage	2.418	1.630	1.343	2.418	1.630	1.343	
Oats Acreage	•974	1.038	1.123	•974	1.038	1.123	
Hay cut Acreage	.870	•844	.781	.87 0	.844	.781	
Tobacco Acreage	**	**	**	**	**	**	
	Region II						
Farms	• 598	*	.713	.596	*	.710	
Farmland	1.775	1.194	1.194	1.335	.992	•992	
Cropland Harvested	3.914	3.174	2 .5 85	4.019	3.135	2.654	
Corn Acreage	3.241	2.776	1.950	2.092	1.749	1.418	
Wheat Acreage	11.142	9.581	8.238	11.046	9.495	8.148	
Cotton Acreage	6.188	2.228	2.017	6.135	2.218	2.003	
Soybean Acreage	20.219	18.021	17.975	20.043	17.863	17.818	
Oats Acreage	12.685	13.320	9.105	8.194	5.794	5.323	
Hay cut Acreage	7.296	6,585	5.405	6.567	5.620	5.304	
Tobacco Acreage	7.837	5,902	6.804	7.769	5.850	6.745	
			Regi	on III			
Farms	.553	*	. 440	•546	*	•436	
Farmland	5.612	. 878	. 873	•721	• 688	. 688	
Cropland Harvested	1.086	. 604	•577	1.098	•555	• 563	
Corn Acreage	1.142	•750	.643	1.092	.740	.641	
Wheat Acreage	1.326	1.186	1.208	1.347	1.134	1.129	
Cotton Acreage	12.289	15.264	11.742	12,182	14.455	12.039	
Soybean Acreage	29.449	31,518	26.694	29.196	31.246	26 •450	
Oats Acreage	2.536	2.025	2.130	2.473	2.018	2.136	
Hay cut Acreage	1.133	.961	.619	1.042	.944	.639	
Tobacco Acreage	**	**	**	**	**	**	

*Date on number of farms not available.

		All Farms	;	Excluding Large Farms			
	Asso	ciation R	ule	Association Rule			
Item	Head-	Closed	Weighted	Head-	Closed	Weighted	
	quarters	Segment	Segment	quarters	Segment	Segment	
			Regi	on I			
Farms	0.716	*	0.798	0.716	*	0.798	
Farmland	•744	•942	•942	•744	•942	•942	
Cropland Harvested	. 828	.960	.966	. 828	. 960	• 966	
Corn Acreage	.810	.839	-922	. 810	•839	•922	
Wheat Acreage	.623	. 857	•905	•623	. 857	•905	
Cotton Acreage	**	**	**	**	**	**	
Soybean Acreage	• 6 86	.705	.793	. 686	•705	•793	
Oats Acreage	•750	.722	.895	•750	•722	. 895	
Hay cut Acreage	•732	.751	•752	•732	•751	•752	
Tobacco Acreage	**	**	**	**	**	**	
	Region II						
Farms	.799	*	.686	.799	*	•6 86	
Farmland	. 887	•932	•932	. 832	.918	.918	
Cropland Harvested	. 851	•947	•936	. 828	•930	. 928	
Corn Acreage	.828	.851	.834	. 657	.704	. 758	
Wheat Acreage	. 608	.731	•656	.610	•730	.657	
Cotton Acreage	070	. 685	•734	- .070	. 405	• 594	
Soybean Acreage	•574	. 820	.702	• 574	. 824	. 699	
Oats Acreage	.697	.769	•723	.182	•374	.360	
Hay cut Acreage	.660	.671	.690	. 360	•379	• 538	
Tobacco Acreage	•962	•957	•960	•962	. 957	•960	
	Region III						
Farms	.691	*	•714	.704	*	.718	
Farmland	. 855	.319	.319	•429	•378	. 378	
Cropland Harvested	• 597	. 817	•730	. 684	•772	. 740	
Corn Acreage	. 430	. 559	.661	.491	• 58 2	•663	
Wheat Acreage	•446	.892	.632	•537	.827	•6 65	
Cotton Acreage	•663	.743	.741	.663	.652	.712	
Soybean Ac rea g e	<u>،</u> 250	• 578	•344	. 250	• 578	.344	
Oats Acreage	013	•572	•576	•571	. 654	•738	
Hay cut Acreage	•388	•388	•479	• 500	•367	. 453	
Tobacco Acreage	**	**	**	**	**	**	

Table 3. Estimated Average Within Primary Sampling Unit <u>Correlations</u> between 1954 and 1959 Area Segment Totals obtained with Three Rules of Associating Farmland with Sample Area Segments, by Region, with and without Farms of 2000 or More Acres.

*Data on number of farms not available.

ments with the same rules of association, are shown in Table 4. As compared with the variances of the simple expansions shown in Table 1, the results are as would be expected from the magnitudes of the correlation coefficients involved. Most items show some improvement in Regions I and II. Improvement was least noticeable in Region III.

The superiority of the Closed and Weighted Segment approaches is quite noticeable and as expected from the higher correlations. In a number of instances, the indicated gain in statistical efficiency would more than offset the increased cost, if any, of the Weighted Segment approach.

Ratio Estimates

The relative variances of 1959 ratio estimates, comparable to the difference estimates of the preceding section, are given in Table 5. As anticipated there are no striking differences between the results in Tables 4 and 5. In some cases the difference estimates are better but in others the ratio estimates have less variability. There seems to be a small edge in favor of ratio estimates.

Regression Estimates

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Regression estimates should have less variability than difference or ratio estimates because sampling fluctuations in base data have only a negligible effect, whereas in difference and ratio estimates such fluctuations exert considerable effect. For the present computations the effects of sampling fluctuations in 1954 base data on the regression estimates were ignored completely. The results are shown in Table 6.

As expected, the relative variances are generally smaller than for the other kinds of estimates. The Closed and Weighted Segment Rules of association again show considerable superiority, with the latter being a bit better. Eliminating the large farms from the sample did not change the relative variances appreciably except in a few instances. This was also true for the difference and ratio estimates. The items affected the most by removing the large farms were oats acreage in Regions II and III and farmland in Region III.

5. Summary and Conclusions

On the basis of the variances observed in this study, the Closed and Weighted Segment Rules of association are decidely preferable to the Headquarters Rule. This holds true for each of the four methods of estimation considered. Relative variances generally are a bit lower for the Weighted Segment Rule than for the Closed Segment Rule. The elimination of farms of 2,000 acres or more from the sample reduced the relative variances somewhat, but the reduction was not particularly striking. However, the number of large farms in the sample was small. When large farms are eliminated from an area sample and treated separately, the relative variance of the estimate for the two strata combined may be reduced appreciably.

When current estimates are computed by difference, ratio or regression procedures with matching data from a previous Census year, an appreciable reduction can be achieved in the relative variances of the agricultural items studied. The reductions in Region III (where the correlations between the 1954 and 1959 data were lower) would probably be less than in Regions I and II. Ratio and difference estimates would tend to have similar precision, with perhaps a slight edge for the ratio estimates. As expected, regression estimates would be the most precise.

The gain in precision to be achieved with the Closed and Weighted Segment Rules as compared to the Headquarters approach is even greater with estimation procedures which make use of base year data, since the correlations are generally higher for these approaches.

Obviously, other considerations in addition to the size of the sampling error must be taken into account when recommending one procedure over another. However, when the farm as a whole is regarded as the unit of observation, the Weighted Segment Rule appears to be preferable to other rules that might be considered for associating farms with sample segments. Not only are sampling errors smaller with this rule; the rule is less likely to be misinterpreted or misapplied by interviewers in the field.

For items where the farm as a whole does not necessarily have to serve as the unit of observation, the Closed Segment Rule has much to commend it. Perhaps the most serious objection to this rule is that some supplemental procedure must be used in conjunction with it to arrive at a count of number of farms and to obtain data on items that apply to the farm as a whole. The rule does have the advantage that interviewers have less difficulty determining the land to be covered in the enumeration. If all of the tracts enumerated are delineated on aerial photos, interviewers and supervisors can perform approximate visual verifications on much of the data reported by respondents. Both of these properties should have the effect of reducing nonsampling errors.

Table 4.Estimated Average Within Primary Sampling Unit Relative Variances (n=1)for Difference Estimatesobtained with Three Rules of Associating FarmLand with Sample Area Segments, by Region, with and without Farms of2000 or More Acres.

	All Farms Association Rule			Exclud	Excluding Large Farms Association Rule			
				Asso				
Item	Head-	Closed	Weighted	Head-	Closed	Weighted		
	quarters	Segment	Segment	quarters	Segment	Segment		
			Reg	<u>ion I</u>				
Farms	0.286	*	0.178	0.286	*	0.178		
Farmland	.35 0	•049	.049	. 350	. 049	.049		
Cropland Harvested	•274	.045	•039	. 274	. 045	.039		
Corn Acreage	.335	.257	.112	.335	. 257	.112		
Wheat Acreage	2,564	1.214	• 584	2.564	1.214	. 584		
Cotton Acreage	**	**	**	**	**	**		
Soybean Acreage	1.775	.909	.640	1.775	.909	.640		
Oats Acreage	.663	.713	• 244	.6 63	.713	•244		
Hay cut Acreage	. 445	•446	.367	.445	•446	.367		
Tobacco Acreage	**	**	**	**	**	**		
			Regi	on II				
Farms	.589	*	.481	•584	*	.476		
Farmland	. 453	.222	.222	• 540	.272	.272		
Cropland Harvested	1.226	.328	.320	1.420	.437	.374		
Corn Acreage	1.213	. 857	•734	1.614	1.088	.848		
Wheat Acreage	7.030	4.473	4.854	6.945	4.445	4.765		
Cotton Acreage	14.314	4.494	2.195	12.353	5.975	2.703		
Soybean Acreage	16.882	11.100	12.383	16.736	11.094	12.393		
Oats Acreage	15.347	13.032	12.215	29,185	30.813	21.817		
Hay cut Acreage	4.197	3.620	2.957	6.690	5.454	4.280		
Tobacco Acreage	1.742	1.816	1.322	1.727	1.800	1.310		
	Region III							
Farms	.391	*	.309	•374	*	.296		
Farmland	1.514	2,551	2.551	.974	.716	.716		
Cropland H arveste d	1.309	.262	.367	1.051	.305	.322		
Corn Acreage	1.120	•574	.419	1.010	• 548	•42 3		
Wheat Acreage	4.029	. 486	.780	2.546	.591	.754		
Cotton Acreage	48.395	13.032	14.932	47.972	21.355	17.584		
Soybean Acreage	29,054	21.442	24.622	28,804	21,258	24.397		
Oats Acreage	8,296	2.174	1.799	3.093	2.045	1.189		
Hay cut Acreage	2.114	1.544	.924	1.587	1.573	.974		
Tobacco Acreage	**	**	**	**	**	**		

*Data on number of farms not available.

Table 5. Estimated Average Within Primary Sampling Unit Relative Variances (n=1) for <u>Ratio Estimates</u> obtained with Three Rules of Associating Farm Land with Sample Area Segments, by Region, with and without Farms of 2000 or hore Acres.

	All Farms Association Rule			Excluding Large Farms		
				Association Rule		
Iten	Head-	Closed	Weighted	Head-	Closed	Weighted
	quarters	Segment	Segment	quarters	Segment	Segment
			Re	gion I		
Farms	0,209	*	0.140	0,209	*	0.140
Farmland	.309	.048	.048	.309	.048	•043
Cropland Harvested	.278	.045	.038	.278	.045	.03 8
Corn Acreage	.374	.282	.123	.374	.202	.123
Wheat Acreage	2,198	1.320	.540	2,198	1.320	.540
Cotton Acreage	**	**::	**	**	**	**
Soybean Acreage	2.035	1.206	.825	2.035	1.206	.825
Oats Acreage	.507	.533	.236	.507	• 533	•236
Hay cut Acreage	•454	.393	.348	.454	.393	•348
Tobacco Acreage	**	**	**	tric	**	**
			Re	gion II		
Farms	.257	*	.378	.255	*	.376
Farmland	.394	.162	.162	.453	.176	.176
Cropland Harvested	1.118	400	.331	1.301	.524	.446
Corn Acreage	1.022	.767	600	1.230	.922	.639
Wheat Acreage	7.573	4.953	5,463	7.442	4.899	5.326
Cotton Acreage	9.617	1.514	1.023	9.732	2.828	1,538
Sovbean Acreage	13.564	6.915	9.618	13.447	6.603	9.486
Oats Acreage	5.388	6.379	4.987	16.951	11.466	8,512
Hay cut Acreage	4.163	3.647	2.859	6.295	5.263	3.945
Tobacco Acreage	.602	.517	.540	.596	.512	.535
	Region III					
Farms	.3 02	*	.234	.287	*	.227
Farmland	1.742	2.052	2.052	.787	.705	.705
Cropland Harvested	924	.226	.320	.756	.259	.290
Corn Acreage	1.041	.556	.392	.931	•530	. 397
Wheat Acreage	2,448	.314	.708	1.722	•421	• 684
Cotton Acreage	24.565	7.005	6.383	8.830	9.022	6.531
Soybean Acreage	58,914	42.573.	48.035	58.388	42.174	47.613
Oats Acreage	4.615	1.523	1.470	2.017	1.361	.979
Hay cut Acreage	1.324	1,112	•742	1.014	1.120	•7 86
Tobacco Acreage	**	**	**	**	**	**

*Data on number of farms not available.

Table 6. Estimated Average Within Primary Sampling Unit Relative Variances (n=1) for <u>Regression Estimates</u> obtained with Three Rules of Associating Farm Land with Sample Area Segments, by Region, with and without Farms of 2000 or More Acres.

	All Farms			Exclud	Excluding Large Farms		
	Asso	ciation P.	ule	<u>Asso</u>	ule		
Item	Head-	Closed	Weighted	Head-	Closed	Weighted	
	quarters	Segment	Segment	quarters	Segment	Segment	
			Re	egion I			
Farms	0.189	*	0,138	0,189	*	0.138	
Farmland	.243	•048	.048	.243	.048	.043	
Cropland Harvested	.226	.044	. C38	.226	.044	.038	
Corn Acreage	.314	.254	.112	.314	.254	.112	
Wheat Acreage	1,315	. 690	.361	1.315	.690	.361	
Cotton Acreage	**	**	**	**	**	**	
Soybean Acreage	1.281	.319	.499	1.281	.819	.499	
Oats Acreage	•426	.497	•224	•425	.497	.224	
Hay cut Acreage	.403	.36 8	.339	.403	.368	.339	
Tobacco Acreage	**	**	**	**	**	**	
	Region II						
Farms	.216	**	.378	.215	*	.376	
Farmland	.378	. 157	.157	.410	.157	. 157	
Cropland Harvested	1.081	.326	.318	1.264	.431	.367	
Corn Acreage	1.020	.483	.594	1.189	.832	.603	
Wheat Acreage	7.025	4.465	4.689	6.936	4.440	4.626	
Cotton Acreage	6.153	1,181	.932	6,105	1.854	1.296	
Soybean Acreage	13,563	5,918	9.129	13.445	5.743	9.104	
Oats Acreage	6.515	5.441	4.236	7.922	4.986	4.632	
Hay cut Acreage	4.121	3.619	2.831	5.718	4.814	3.772	
Tobacco Acreage	.592	.500	.532	.586	.495	.528	
	Region III						
Farms	.289	*	.215	.275	*	.212	
Farmland	1.511	.789	.789	.589	. 590	590	
Cropland Harvested	. 699	.201	.270	.585	.224	.255	
Corn Acreage	.931	.516	.362	.829	490	.360	
Wheat Acreage	1.062	242	.646	959	358	.630	
Cotton Acreage	6.891	6.832	5.300	6.831	8,303	5.933	
Soybean Acreage	27.514	20,973	23.534	27.377	20.791	23.319	
Oats Acreage	2.535	1.363	1.424	1.666	1,155	.971	
Hay cut Acreage	.963	.317	.477	.781	.317	.508	
Tobacco Acreage	**	**	**	**	***	**	

*Data on number of farms not available.